ECE541: Stochastic Signals and Systems

Fall 2018

Chapter 7: Convergence of Random Sequences Dr. Salim El Rouayheb

Scribe: Abhay Ashutosh Donel, Qinbo Zhang, Peiwen Tian, Pengzhe Wang, Lu Liu

## 1 Random sequence

**Definition 1.** An infinite sequence  $X_n$ , n = 1, 2, ..., of random variables is called a random sequence.

## 2 Convergence of a random sequence

Example 1. Consider the sequence of real numbers

$$X_n = \frac{n}{n+1}, \ n = 0, 1, 2, \dots$$

This sequence converges to the limit l = 1. We write

$$\lim_{n \to \infty} X_n = l = 1.$$

This means that in any neighbourhood around 1 we can trap the sequence, i.e.,

$$\forall \epsilon > 0, \quad \exists n_0(\epsilon) \quad s.t. \quad for \ n \ge n_0(\epsilon) \quad |X_n - l| \le \epsilon.$$

We can pick  $\epsilon$  to be very small and make sure that the sequence will be trapped after reaching  $n_0(\epsilon)$ . Therefore as  $\epsilon$  decreases  $n_0(\epsilon)$  will increase. For example, in the considered sequence:

$$\epsilon = \frac{1}{2}, \qquad n_0(\epsilon) = 2,$$
  
 $\epsilon = \frac{1}{1000}, \qquad n_0(\epsilon) = 1001.$ 

### 2.1 Almost sure convergence

**Definition 2.** A random sequence  $X_n$ , n = 0, 1, 2, 3, ..., converges almost surely, or with probability one, to the random variable X iff

$$P(\lim_{n \to \infty} X_n = X) = 1.$$

We write

 $X_n \xrightarrow{a.s.} X.$ 

**Example 2.** Let  $\omega$  be a random variable that is uniformly distributed on [0, 1]. Define the random sequence  $X_n$  as  $X_n = \omega^n$ .

So 
$$X_0 = 1$$
,  $X_1 = \omega$ ,  $X_2 = \omega^2$ ,  $X_3 = \omega^3$ ,...

Let us take specific values of  $\omega$ . For instance, if  $\omega = \frac{1}{2}$ 

$$X_0 = 1, \ X_1 = \frac{1}{2}, \ X_2 = \frac{1}{4}, \ X_3 = \frac{1}{8}, \dots$$

We can think of it as an urn containing sequences, and at each time we draw a value of  $\omega$ , we get a sequence of fixed numbers. In the example of tossing a coin, the output will be either heads or tails. Whereas, in this case the output of the experiment is a random sequence, i.e., each outcome is a sequence of infinite numbers.

**Question:** Does this sequence of random variables converge?

**Answer:** This sequence converges to

$$X = \begin{cases} 0 & \text{if } \omega \neq 1 \text{ with probability } 1 = P(\omega \neq 1) \\ 1 & \text{if } \omega = 1 \text{ with probability } 0 = P(\omega = 1) \end{cases}$$

Since the pdf is continuous, the probability  $P(\omega = a) = 0$  for any constant a. Notice that the convergence of the sequence to 1 is possible but happens with probability 0.

Therefore, we say that  $X_n$  converges almost surely to 0, i.e.,  $X_n \xrightarrow{a.s.} 0$ .

**Example 3.** Consider a random variable  $\omega \in \Omega = [0,1]$  uniformly distributed on [a,b],  $0 \le a \le b \le 1$ , and the sequence  $X_n(\omega)$ , n = 1, 2, ..., defined by:

$$X_n(\omega) = \begin{cases} 1 & \text{if } 0 \le \omega < \frac{n+1}{2n}, \\ 0 & \text{otherwise.} \end{cases}$$

Also, define the random variable X defined by:

$$X(\omega) = \begin{cases} 1 & if \ 0 \le \omega < \frac{1}{2} \\ 0 & otherwise. \end{cases}$$

Show that  $X_n \xrightarrow{a.s.} X$ .

**Solution:** Define the set A as follows:

$$A = \{ \omega \in \Omega : \lim_{n \to +\infty} X_n(\omega) = X(\omega) \}.$$

We need to prove that P(A) = 1. Let's first find A. Note that  $\frac{n+1}{2n} > \frac{1}{2}$ , so for any  $\omega \in [0, \frac{1}{2}]$ , we have

$$X_n(\omega) = X(\omega) = 1.$$

Therefore, we conclude that  $[0, 0.5 \subset A$ . Now, if  $\omega > \frac{1}{2}$ , then

$$X(\omega) = 0.$$

Also, since  $2\omega - 1 > 0$ , we can write

$$X_n(\omega) = 0, \quad \forall n > \frac{1}{2\omega - 1}.$$

Therefore,

$$\lim_{n \to +\infty} X_n(\omega) = X(\omega) = 0, \quad \forall \omega > \frac{1}{2}.$$

We conclude  $[0.5, 1] \subset \Omega$ . You can check that  $\omega = 0.5 \notin A$ , since

$$X_n(0.5) = 1, \quad \forall n,$$

while X(0.5) = 0. We conclude

$$A = \left[0, \frac{1}{2} \left[\cup\right] \frac{1}{2}, 1\right] = \Omega - \left\{\frac{1}{2}\right\}.$$

Since P(A) = 1, we conclude  $X_n \xrightarrow{a.s.} X$ .

**Theorem 1.** Consider the sequence  $X_1, X_2, X_3, \ldots$  For any  $\epsilon > 0$ , define the set of events

$$A_m = \{ |X_n - X| < \epsilon, \ \forall n \ge m \}$$

Then  $X_n \xrightarrow{a.s.} X$  if and only if for any  $\epsilon > 0$ , we have

$$\lim_{m \to +\infty} P(A_m) = 1.$$

**Example 4.** Let  $X_1, X_2, X_3, \ldots$  be independent random variables, where  $X_n \sim Bernoulli\left(\frac{1}{n}\right)$  for  $n = 2, 3, \ldots$  The goal here is to check whether  $X_n \xrightarrow{a.s.} 0$ .

- 1. Check that  $\sum_{n=1}^{+\infty} P(|X_n| > \epsilon) = +\infty$ .
- 2. Show that the sequence  $X_1, X_2, \ldots$  does not converge to 0 almost surely using Theorem 1.

#### Solution:

1. We first note that for  $0 < \epsilon < 1$ , we have

$$\sum_{n=1}^{+\infty} P(|X_n| > \epsilon) = \sum_{n=1}^{+\infty} P(|X_n| > \epsilon) = \sum_{n=1}^{+\infty} \frac{1}{n} = +\infty.$$

2. To use Theorem 1, we define

$$A_m = \{ |X_n| < \epsilon, \ \forall n \ge m \}.$$

Note that for  $0 < \epsilon < 1$ , we have

$$A_m = \{X_n = 0, \ \forall n \ge m\}.$$

According to Theorem 1, it suffices to show that

$$\lim_{m \to +\infty} P(A_m) < 1.$$

We can in fact show that  $\lim_{\to +\infty} P(A_m) = 0$ . To show this, we will prove  $P(A_m) = 0$ , for every  $m \ge 2$ . For  $0 < \epsilon < 1$ , we have

$$P(A_m) = P\left(\{X_n = 0, \forall n \ge m\}\right)$$
  

$$\leq P\left(\{X_n = 0, \forall n = m, m+1, \dots, N\}\right) \text{ (for every positive integer } N \ge m)$$
  

$$= P(X_m = 0)P(X_{m+1} = 0)\dots P(X_N = 0) \text{ (since the } X'_i \text{s are independent})$$
  

$$= \frac{m-1}{m} \cdot \frac{m}{m+1} \dots \frac{N-1}{N}$$
  

$$= \frac{m-1}{N}.$$

Thus, by choosing N large enough, we can show that  $P(A_m)$  is less than any positive number. Therefore,  $P(A_m) = 0$ , for all  $m \ge 2$ . We conclude that  $\lim_{m\to+\infty} P(A_m) = 0$ . Thus, according to Theorem 1, the sequence  $X_1, X_2, \ldots$  does not converge to 0 almost surely.

**Theorem 2.** Strong law of large numbers

Let  $X_1, X_2, X_3, \ldots, X_i$  be iid random variables.  $E[X_i] = \mu, \forall i$ . Let

$$S_n = \frac{X_1 + X_2 + \ldots + X_n}{n}.$$

Then

$$P\left[\lim_{n \to \infty} |S_n - \mu| \ge \epsilon\right] = 0.$$

Using the language of this chapter:

$$S_n \xrightarrow{a.s.} \mu.$$

#### 2.2 Convergence in probability

**Definition 3.** A random sequence  $X_n$  converges to the random variable X in probability if

$$\forall \epsilon > 0 \quad \lim_{n \to \infty} \Pr\{|X_n - X| \ge \epsilon\} = 0.$$

We write :

 $X_n \xrightarrow{p} X.$ 

**Example 5.** Consider a random variable  $\omega$  uniformly distributed on [0,1] and the sequence  $X_n$  given in Figure ??. Notice that only  $X_2$  or  $X_3$  can be equal to 1 for the same value of  $\omega$ . Similarly, only one of  $X_4, X_5, X_6$  and  $X_7$  can be equal to 1 for the same value of  $\omega$  and so on and so forth.

**Question:** Does this sequence converge?

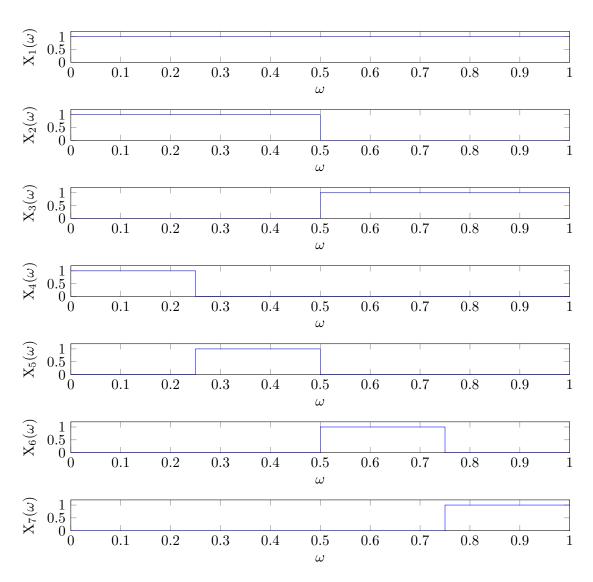


Figure 1: Plot of the distribution of  $X_n(\omega)$ 

**Answer:** Intuitively, the sequence will converge to 0. Let us take some examples to see how the sequence behave.

for 
$$\omega = 0$$
:  
 $\lim_{n=1}^{1} \lim_{n=2}^{1000} \frac{10000000}{n=4} \dots$   
for  $\omega = \frac{1}{3}$ :  
 $\lim_{n=1}^{1} \lim_{n=2}^{100} \frac{010000100000}{n=4} \dots$ 

From a calculus point of view, these sequences never converge to zero because there is always a "jump" showing up no matter how many zeros are preceding (Fig. ??); for any  $\omega : X_n(\omega)$  does not converge in the "calculus" sense. Which means also that  $X_n$  does not converge to zero almost surely (a.s.).

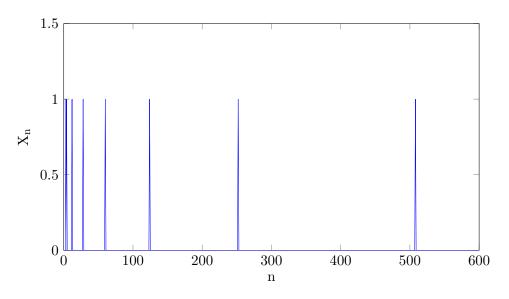


Figure 2: Plot of the sequence for  $\omega = 0$ 

This sequence converges in probability since

$$\lim_{n \to \infty} P\left(|X_n - 0| \ge 0\right) = 0 \quad \forall \epsilon > 0$$

**Remark 1.** The observed sequence may not converge in "calculus" sense because of the intermittent "jumps"; however the frequency of those "jumps" goes to zero when n goes to infinity.

**Example 6.** Consider a random variable  $\omega$  uniformly distributed over [0,1], and the sequence  $X_n(\omega)$  defined as:

$$X_n(\omega) = \begin{cases} 1 & \text{for } \omega \leq \frac{1}{n} \\ 0 & \text{otherwise} \end{cases}$$

Question: Does this sequence converge a.s.? in probability?

### Solution:

1. First, we will use Theorem 1 to show that the sequence does not converge a.s.. Let

$$A_m = \{ |X_n| < \epsilon, \ \forall n \ge m \}.$$

Note that for  $0 < \epsilon < 1$ , we have

$$A_m = \{X_n = 0, \ \forall n \ge m\}.$$

$$\begin{split} P(A_m) &= P\left(\{X_n = 0, \ \forall n \ge m\}\right) \\ &\leq P\left(\{X_n = 0, \ \forall n = m, m+1, \dots, N\}\right) \text{ (for every positive integer } N \ge m) \\ &= P(X_m = 0)P(X_{m+1} = 0) \dots P(X_N = 0) \text{ (since the } X'_is \text{ are independent)} \\ &= P(w > \frac{1}{m})P(w > \frac{1}{m+1}) \dots P(w > \frac{1}{N}) \\ &= \frac{m-1}{m} \cdot \frac{m}{m+1} \dots \frac{N-1}{N} \\ &= \frac{m-1}{N}. \end{split}$$

We conclude that  $\lim_{m\to+\infty} P(A_m) = 0$ . Thus, according to Theorem 1, the sequence  $X_1, X_2, \ldots$  does not converge to 0 almost surely.

2. Now we check for convergence in probability.

$$Pr(X_n \ge \epsilon) = Pr(X_n = 1) = Pr(w \le \frac{1}{n}) = \frac{1}{n}.$$

Hence,

$$\lim_{n \to +\infty} \Pr(X_n \ge \epsilon) = \lim_{n \to +\infty} \frac{1}{n} = 0.$$

Therefore,  $X_n \xrightarrow{p.} 0$ .

**Theorem 3.** Weak law of large numbers

Let  $X_1, X_2, X_3, \ldots, X_i$  be iid random variables.  $E[X_i] = \mu, \forall i$ . Let

$$S_n = \frac{X_1 + X_2 + \ldots + X_n}{n}$$

Then

$$P\left[|S_n - \mu| \ge \epsilon\right] \xrightarrow[n \to \infty]{} 0.$$

Using the language of this chapter:

 $S_n \xrightarrow{p_{\cdot}} \mu.$ 

### 2.3 Convergence in mean square

**Definition 4.** A random sequence  $X_n$  converges to a random variable X in mean square sense if

$$\lim_{n \to \infty} E\left[\left|X - X_n\right|^2\right] = 0.$$

We write:

$$X_n \xrightarrow{m.s.} X.$$

**Remark 2.** In mean square convergence, not only the frequency of the "jumps" goes to zero when n goes to infinity; but also the "energy" in the jump should go to zero.

**Example 6.** (Revisited) Does  $X_n$  converge in m.s.?

Answer:

$$E\left[|X_n - 0|^2\right] = 1 \cdot P(w \le \frac{1}{n}) + 0 \cdot P(w > \frac{1}{n}) = \frac{1}{n}$$
$$\lim_{n \to \infty} E\left[|X_n - 0|^2\right] = \lim_{n \to \infty} \frac{1}{n} = 0.$$

Therefore,  $X_n \xrightarrow{m.s.} 0$ .

In the next example, we replace 1 by  $\sqrt{n}$  in Example 5.

**Example 7.** Consider a random variable  $\omega$  uniformly distributed over [0,1], and the sequence  $X_n(\omega)$  defined as:

$$X_n(\omega) = \begin{cases} \sqrt{n} & \text{for } \omega \leq \frac{1}{n} \\ 0 & \text{otherwise} \end{cases}$$

Note that  $P(X_n = a_n) = \frac{1}{n}$  and  $P(X_n = 0) = 1 - \frac{1}{n}$ .

Question: Does this sequence converge a.s.? in probability? in m.s.?

#### Answer:

- 1. Almost sure convergence:  $X_n$  does not converge a.s. for the same reasons as Example 5.
- 2. Convergence in probability:  $X_n \xrightarrow{p} 0$  for the same reasons as Example 5. Namely,

$$\lim_{n \to +\infty} \Pr(X_n \ge \epsilon) = \lim_{n \to +\infty} \Pr(X_n = \sqrt{n}) = \lim_{n \to +\infty} \frac{1}{n} = 0$$

(Flash Forward: almost sure convergence  $\Rightarrow$  convergence in probability, but convergence in probability  $\Rightarrow$  almost sure convergence.)

3. Mean Square Convergence:

$$E\left[\left|X_{n}-0\right|^{2}\right] = n \cdot P\left(w \leq \frac{1}{n}\right) + 0 \cdot P\left(w > \frac{1}{n}\right) = n \cdot \frac{1}{n} = 1.$$

Hence,

 $\lim_{n \to \infty} E\left[|X_n - 0|^2\right] = 1 \Rightarrow X_n \text{ does not converge in m.s. to } 0.$ 

### 2.4 Convergence in distribution

**Definition 5.** (First attempt) A random sequence  $X_n$  converges to X in distribution if when n goes to infinity, the values of the sequence are distributed according to a known distribution. We say

 $X_n \xrightarrow{d_{\cdot}} X_{\cdot}$ 

**Example 8.** Consider the sequence  $X_n$  defined as:

$$X_n = \begin{cases} X_i \sim B(\frac{1}{2}) & \text{for } i = 1\\ (X_{i-1} + 1) \mod 2 = X \oplus 1 & \text{for } i > 1 \end{cases}$$

**Question:** In which sense, if any, does this sequence converge?

**Answer:** This sequence has two outcomes depending on the value of  $X_1$ :

$$X_1 = 1, \quad X_n : 101010101010...$$
  
 $X_1 = 0, \quad X_n : 010101010101...$ 

- 1. Almost sure convergence:  $X_n$  does not converge almost surely because the probability of every jump is always equal to  $\frac{1}{2}$ .
- 2. Convergence in probability:  $X_n$  does not converge in probability because the frequency of the jumps is constant equal to  $\frac{1}{2}$ .
- 3. Convergence in mean square:  $X_n$  does not converge to  $\frac{1}{2}$  in mean square sense because

$$\lim_{n \to \infty} E\left[|X_n - \frac{1}{2}|^2\right] = E\left[X_n^2 - X_n + \frac{1}{4}\right],$$
$$= E[X_n^2] - E[X_n] + \frac{1}{4}$$
$$= 1^2 \frac{1}{2} + 0^2 \frac{1}{2} - 0 + \frac{1}{4},$$
$$= \frac{1}{2}.$$

4. Convergence in distribution: At infinity, since we do not know the value of  $X_1$ , each value of  $X_n$  can be either 0 or 1 with probability  $\frac{1}{2}$ . Hence, any number  $X_n$  is a random variable  $\sim B(\frac{1}{2})$ . We say,  $X_n$  converges in distribution to Bernoulli $(\frac{1}{2})$  and we denote it by:

$$X_n \xrightarrow{d} Ber(\frac{1}{2}).$$

**Example 9.** (Central Limit Theorem)Consider the zero-mean, unit-variance, independent random variables  $X_1, X_2, \ldots, X_n$  and define the sequence  $S_n$  as follows:

$$S_n = \frac{X_1 + X_2 + \dots + X_n}{\sqrt{n}}$$

The CLT states that  $S_n$  converges in distribution to N(0,1), i.e.,

$$S_n \xrightarrow{d} N(0,1).$$

#### Theorem 4.

 $\left. \begin{array}{l} Almost \ sure \ convergence \\ Convergence \ in \ mean \ square \end{array} \right\} \Rightarrow Convergence \ in \ probability \Rightarrow convergence \ in \ distribution. \end{array} \right\}$ 

Note:

- There is no relation between Almost Sure and Mean Square Convergence.
- The relation is unidirectional, i.e., convergence in distribution does not imply convergence in probability neither almost sure convergence nor mean square convergence.

# 3 Convergence of a random sequence

**Example 1:** Let the random variable U be uniformly distributed on [0, 1]. Consider the sequence defined as:

$$X(n) = \frac{(-1)^n U}{n}.$$

**Question:** Does this sequence converge? if yes, in what sense(s)?

#### Answer:

1. Almost sure convergence: Suppose

$$U = a$$
.

The sequence becomes

$$X_1 = -a,$$
  

$$X_2 = \frac{a}{2},$$
  

$$X_3 = -\frac{a}{3},$$
  

$$X_4 = \frac{a}{4},$$
  

$$\vdots$$

In fact, for any  $a \in [0, 1]$ 

$$\lim_{n \to \infty} X_n = 0,$$

therefore,  $X_n \xrightarrow{a.s.} 0$ .

**Remark 3.**  $X_n \xrightarrow{a.s.} 0$  because, by definition, a random sequence converges almost surely to the random variable X if the sequence of functions  $X_n$  converges for all values of U except for a set of values that has a probability zero.

2. Convergence in probability: Does  $X_n \xrightarrow{p} 0$ ? Recall from theorem 13 of lecture 17:

$$\left. \begin{array}{c} \text{a.s.} \\ \text{m.s.} \end{array} \right\} \Rightarrow \text{p.} \Rightarrow \text{d.}$$

which means that by proving almost-sure convergence, we get directly the convergence in probability and in distribution. However, for completeness we will formally prove that  $X_n$  converges to 0 in probability. To do so, we have to prove that

$$\lim_{n \to \infty} P(|X - 0| \ge \epsilon) = 0 \quad \forall \epsilon > 0,$$
  
$$\Rightarrow \lim_{n \to \infty} P(|X_n| \ge \epsilon) = 0 \quad \forall \epsilon > 0.$$

By definition,

$$|X_n| = \frac{U}{n} \le \frac{1}{n}$$

Thus,

$$\lim_{n \to \infty} P\left(|X_n| \ge \epsilon\right) = \lim_{n \to \infty} P\left(\frac{U}{n} \ge \epsilon\right),\tag{1}$$

$$=\lim_{n\to\infty}P\left(U\ge n\epsilon\right),\tag{2}$$

$$=0.$$
 (3)

Where equation 3 follows from the fact that finding  $U \in [0, 1]$ .

3. Convergence in mean square sense: Does  $X_n$  converge to 0 in the mean square sense?

In order to answer this question, we need to prove that

$$\lim_{n \to \infty} E\left[|X_n - 0|^2\right] = 0.$$

We know that,

$$\lim_{n \to \infty} E\left[|X_n - 0|^2\right] = \lim_{n \to \infty} E\left[X_n^2\right],$$
$$= \lim_{n \to \infty} E\left[\frac{U^2}{n^2}\right],$$
$$= \lim_{n \to \infty} \frac{1}{n^2} E\left[U^2\right],$$
$$= \lim_{n \to \infty} \frac{1}{n^2} \int_0^1 u^2 du,$$
$$= \lim_{n \to \infty} \frac{1}{n^2} \frac{u^3}{3} \Big]_0^1,$$
$$= \lim_{n \to \infty} \frac{1}{3n^2},$$
$$= 0.$$

Hence,  $X_n \xrightarrow{m.s.} 0$ .

4. Convergence in distribution: Does  $X_n$  converge to 0 in distribution? The formal definition of convergence in distribution is the following:

$$X_n \xrightarrow{d.} X \Rightarrow \lim_{n \to \infty} F_{X_n}(x) = F_X(x).$$

Hereafter, we want to prove that  $X_n \xrightarrow{d} 0$ .

Recall that the limit r.v. X is the constant 0 and therefore has the following CDF : Since  $X_n = \frac{(-1)^n U}{n}$ , the distribution of the  $X_i$  can be derived as following:

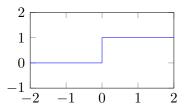
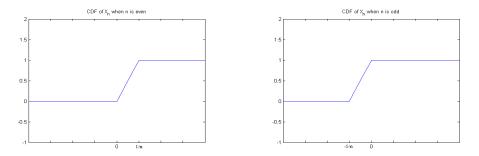


Figure 3: Plot of the CDF of 0



**Remark 4.** At 0 the CDF of  $X_n$  will be flip-flopping between 0 (if n is even) and 1 (if n is odd) (c.f. figure 2) which implies that there is a discontinuity at that point. Therefore, we say that  $X_n$  converges in distribution to a CDF  $F_X(x)$  except at points where  $F_X(x)$  is not continuous.

**Definition 6.**  $X_n$  converges to X in distribution, i.e.,  $X[n] \xrightarrow{d.} X$  iff

 $\lim_{n \to \infty} F_{X_n}(x) = F_X(x) \quad \text{except at points where } F_X(x) \text{ is not continuous.}$ 

**Remark 5.** It is clear here that

$$\lim_{n \to \infty} F_{X_n}(x) = F_x(x) \quad except for \ x = 0.$$

Therefore,  $X_n$  converges to X in distribution. We could have deduced this directly from convergence in mean square sense or almost sure convergence.

**Theorem 5.** a) If  $X_n \xrightarrow{a.s.} X \Rightarrow X_n \xrightarrow{p.} X$ .

- b) If  $X_n \xrightarrow{m.s.} X \Rightarrow X_n \xrightarrow{p.} X$ .
- c) If  $X_n \xrightarrow{p} X \Rightarrow X_n \xrightarrow{d} X$ .
- d) If  $P\{|X_n| \leq Y\} = 1$  for all n for a random variable Y with  $E[Y^2] < \infty$ , then

$$X_n \xrightarrow{p} X \Rightarrow X_n \xrightarrow{m.s} X.$$

*Proof.* The proof is omitted.

**Remark 6.** Convergence in probability allows the sequence, at  $\infty$ , to deviate from the mean for any value with a small probability; whereas, convergence in mean square limits the amplitude of this deviation when  $n \to \infty$ . (We can think of it as energy  $\Rightarrow$  we can not allow a big deviation from the mean).

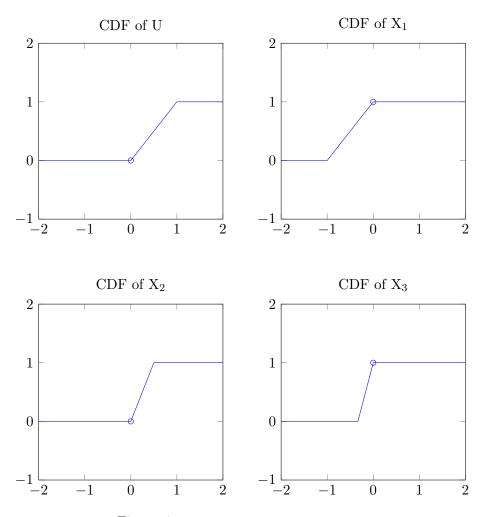


Figure 4: Plot of the CDF of  $U, X_1, X_2$  and  $X_3$ 

# 4 Back to real analysis

**Definition 7.** A sequence  $(x_n)_{n\geq 1}$  is Cauchy if for every  $\epsilon$ , there exists a large number N s.t.

$$\forall m, n > N, |x_m - x_n| < \epsilon \quad \Leftrightarrow \quad \lim_{n, m \to \infty} |x_m - x_n| = 0.$$

Claim 1. Every Cauchy sequence is convergent.

**Counter example 1.** Consider the sequence  $X_n \in \mathbb{Q}$  defined as  $x_0 = 1$ ,  $x_{n+1} = \frac{x_n + \frac{2}{x_n}}{2}$ . The limit of this sequence is given by:

$$l = \frac{l + \frac{2}{l}}{2},$$
  
$$2l^2 = l^2 + 2,$$
  
$$l = \pm\sqrt{2} \notin \mathbb{Q}.$$

This implies that the sequence does not converge in  $\mathbb{Q}$ .

**Counter example 2.** Consider the sequence  $x_n = 1/n$  in (0,1). Obviously it does not converge in (0,1) since the limit  $l = 1 \notin (0,1)$ .

**Definition 8.** A space where every sequence converges is called a complete space.

**Theorem 6.**  $\mathbb{R}$  is a complete space.

*Proof.* The proof is omitted.

**Theorem 7.** Cauchy criteria for convergence of a random sequence.

a) 
$$X_n \xrightarrow{a.s.} X \iff P\left[\lim_{m,n\to\infty} |x_m - x_n| = 0\right] = 1.$$
  
b)  $X_n \xrightarrow{m.s.} X \iff \lim_{m,n\to\infty} E\left[|x_m - x_n|^2\right] = 0.$   
c)  $X_n \xrightarrow{p.} X \iff \lim_{m,n\to\infty} P\left[|x_m - x_n| \ge \varepsilon\right] = 0 \quad \forall \epsilon.$ 

*Proof.* The proofs are omitted.

Example 10. Consider the sequence of example 11 from last lecture,

$$X_n = \begin{cases} X_i \sim B(\frac{1}{2}) & \text{for } i = 1\\ (X_{i-1} + 1) \mod 2 = X \oplus 1 & \text{for } i > 1 \end{cases}$$

**Goal:** Our goal is to prove that this sequence does not converge in mean square using Cauchy criteria.

This sequence has two outcomes depending on the value of  $X_1$ :

$$X_1 = 1, \quad X_n : 101010101010...$$
  
 $X_1 = 0, \quad X_n : 010101010101...$ 

Therefore,

$$E\left[|X_n - X_m|^2\right] = E\left[X_n^2\right] + E\left[X_m^2\right] - 2E\left[X_m X_n\right],$$
  
=  $\frac{1}{2} + \frac{1}{2} - 2E\left[X_m X_n\right].$ 

Consider, without loss of generality, that m > n

$$E[X_n X_m] = \begin{cases} E[X_n X_m] = 0 & \text{if } m - n \text{ is odd,} \\ E[X_n^2] = \frac{1}{2} & \text{if } m - n \text{ is even.} \end{cases}$$

Hence,

$$\lim_{n,m\to\infty} E\left[|X_n - X_m|^2\right] = \begin{cases} 1 & \text{if } m - n \text{ is odd,} \\ 0 & \text{if } m - n \text{ is even,} \end{cases}$$

which implies that  $X_n$  does not converge in mean square by theorem 7-b).

**Lemma 1.** Let  $X_n$  be a random sequence with  $E[X_n^2] < \infty \ \forall n$ .

$$X_n \xrightarrow{m.s.} X$$
 iff  $\lim_{m,n\to\infty} E[X_m X_n]$  exists and is finite.

## Theorem 8. Central limit theorem

Let  $X_1, X_2, X_3, \ldots, X_i$  be iid random variables.  $E[X_i] = 0, \forall i.$  Let

$$Z_n = \frac{X_1 + X_2 + \dots + X_n}{\sqrt{n}}.$$

Then

$$P[Z_n \le z] = \int_{-\infty}^z \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2}{2}} dz.$$

Using the language of this chapter:

$$Z_n \xrightarrow{d.} N(0,1).$$

.